Quantifying the Tradeoff Between Precaution and Yield in the U.S. Sea Scallop Fishery

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Abstract

Fishery reference points in the U.S. sea scallop fishery are set using yield 2 per recruit analysis. Because of uncertainties in the parameters used in this 3 analysis, the estimated reference points are uncertain. For this reason, it Δ is often argued that target fishing mortality rates should be less than the 5 calculated reference points in order to reduce the risk of overfishing. However, 6 precautionary management also can reduce yield by fishing at suboptimal 7 rates. Here, I use Monte-Carlo simulations to quantify the tradeoff between 8 overfishing risk and loss in yield per recruit. At fishing mortalities near F_{MAX} , g the fishing mortality where maximum yield per recruit is obtained, reducing 10 fishing mortality obtains a substantial reduction in the risk of overfishing at 11 little cost of lost yield per recruit. At lower fishing mortality rates, however, 12 the marginal benefit in terms of reduced fishing mortality risk from further 13 reductions in fishing mortality becomes less, and the cost in reduced yield 14 per recruit becomes greater. If implementation uncertainty is added to the 15 analysis, the risk of overfishing as well the loss of yield per recruit is increased, 16 except at F_{MAX} . 17

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18 Introduction

Fishery reference points are uncertain because the models that generate them 19 depend on parameters that are themselves uncertain. For this reason, it has 20 long been recommended that reference points be set on a precautionary basis, 21 so as to minimize the risk of overfishing. This approach has been codified into 22 U.S. law in 1996 and 2006 by revisions to the Magnuson-Stevens fishery act. 23 However, reducing fishing mortality below F_{MSY} will, by definition, reduce the 24 expected yields that can be obtained from the fishery. While precaution gives 25 benefits in that it reduces the risk of overfishing and its concomitant impacts 26 on the marine ecosystem, it also has a cost in that it reduces expected yield. 27 The purpose of this paper is to explore these tradeoffs in setting reference 28 points for the U.S. sea scallop, *Placopecten magellanicus*, fishery. 29

Because stock-recruit relationships for sea scallops are not well defined (and are presumably saturated at current and future biomass levels), reference points for sea scallops have been set using yield per recruit analysis, using F_{MAX} as a proxy for F_{MSY} . The most recent sea scallop stock assessment (NEFSC 2007) estimated $F_{\text{MAX}} = 0.24$ on Georges Bank, $F_{\text{MAX}} = 0.36$ in the Mid-Atlantic, and $F_{\text{MAX}} = 0.29$ for the fishery overall.

Uncertainties in yield per recruit analysis can be assessed by estimating a 36 probability distribution for each of the input parameters and then repeatedly 37 drawing parameters at random from these distributions and performing yield 38 per recruit analysis using these choices (Restrepo and Fox 1988). By repeat-39 ing this procedure a large number of times, the probability distribution of 40 F_{MAX} and the expected yield per recruit at a given fishing mortality can be 41 estimated. From this, the probability of overfishing at a fishing mortality F42 as well as the loss in yield per recruit incurred by fishing at F rather than 43 F_{MAX} can be calculated. 44

Besides the uncertainties in the reference points, there is implementation
error in that the fishing mortality target intended by managers may not be

realized precisely, and the actual fishing mortality may be greater or less
than that intended by management. The effect of such errors will also be
discussed here.

50 Methods

51 Monte-Carlo yield per recruit analysis

A description of basic length-based yield per recruit model used in this analysis can be found in Hart (2003). The yield per recruit calculations depend
on a number of parameters which each carry a level of uncertainty:

⁵⁵ (1) Von Bertalanffy growth parameters K and L_{∞}

 $_{56}$ (2) Shell height/meat weight parameters a and b

- 57 (3) Natural mortality rate M
- 58 (4) Fishery selectivity parameters α and β
- ⁵⁹ (5) The cull size of the catch and the fraction of discards that survive
- 60 (6) The level of incidental fishing mortality, i.e., non-catch mortality caused
- 61 by fishing.
- 62

Each of these parameters were assigned a probability distribution reflect-63 ing their level of uncertainty, as discussed below. For each iteration, choices 64 for each of these parameters were drawn from their distributions, and then 65 a yield per recruit analysis was performed. This was repeated for n = 1000066 iterations for both regions (Georges Bank and Mid-Atlantic) and the results 67 collected. Of particular interest were the expected yield per recruit at a given 68 fishing mortality F and the probability that overfishing would be occurring 69 if fishing mortality was F. The expected yield per recruit was calculated 70 simply as the average of the yield per recruit of each run. The probability of 71 overfishing was estimated as the number of runs for which $F_{\text{MAX}} < F$ divided 72 by the total number of runs. 73

The estimates of three sets of these parameters (K and L_{∞} , a and b, and α and β) are confounded, as reflected by a strong correlation between the estimates. For example, a growth curve with a given K and L_{∞} resembles one with a slightly smaller K and larger L_{∞} , implying a negative correlation between the estimates of the two parameters. In these cases, each parameter pair was simulated as correlated normals. In other cases, gamma distributions were used.

The analyses were done separately in each area (Georges Bank and Mid-81 Atlantic). Expected yields were combined assuming that each area is equally 82 productive. This is approximately correct over the last 25 years, though 83 Georges Bank was more productive over a longer time period, and the Mid-84 Atlantic more productive in recent years. Calculating the probability of 85 overfishing of the combined resource requires additional assumptions regard-86 ing the correlation of parameters in the two regions. It would seem likely 87 that a positive correlation exists, e.g., if the natural mortality estimate of 88 0.1 was underestimated in one region, it is likely that it is also in the other. 89 For that reason, it is assumed here that the corresponding parameters in 90 the two regions are correlated with a correlation of 1. If this correlation is 91 smaller, the variability between the regions would partially cancel, and the 92 probability of overfishing would be somewhat less than calculated here. 93

⁹⁴ Probability distributions for the simulated parameters

The mean, standard error and correlation (when applicable) for each of the simulated parameters is given in Table 1. These estimates were taken from the latest sea scallop stock assessment (NEFSC 2007) or from the literature. When standard errors were not available, they were estimated using reasonable judgement. Details on each of these parameters is given below.

Growth parameters K and L_{∞} . These parameters were estimated using a linear mixed-effects model based on the reading of sea scallop rings from shells collected during the 2001-2006 NEFSC sea scallop surveys (NEFSC 2007). These estimates were recently revised by using a slightly refined model and one additional year of data (Hart and Chute 2009). In order to conform to the NEFSC (2007) reference points, the growth parameters estimated there were used, rather than the updated ones. The difference between these is in any case minimal.

As discussed above, K and L_{∞} were simulated as negatively correlated 108 normals, with their mean, variance and covariance as estimated in NEFSC 109 (2007). The standard errors of K and L_{∞} are very small due to the large 110 amount of data available. The true uncertainty may be greater than this 111 "statistical uncertainty" because of model uncertainties. For example, von 112 Bertalanffy growth appears to well approximate sea scallop growth, but is 113 probably not exactly correct. Such uncertainties are not reflected in the 114 standard errors of the parameters. However, simulations indicate that the 115 mixed-effects model is robust to a number of uncertainties, and likely esti-116 mates the mean growth parameters to within 1% of its true value (Hart and 11 Chute 2009). 118

¹¹⁹ Shell height/meat weight relationships. Meat weight W at shell height ¹²⁰ H is calculated using a formula of the form:

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$$W = \exp(a + b\ln(H)) \tag{1}$$

The parameters a and b were estimated during the last sea scallop bench-122 mark assessment (NEFSC 2007) using a generalized mixed-effects model 123 (GLMM) based on data collected during the 2001-2006 NEFSC annual sea 124 scallop surveys. This analysis was used to obtain estimates of means, vari-125 ances, and covariances of the parameters (Table 1). Similar to the growth 126 parameters, the estimates of a and b are somewhat confounded, so that they 127 have a strong negative correlation. This means that the predicted meat 128 weight at a given shell height carries less uncertainty than it would appear 129

¹³⁰ from the variances of the individual parameters.

Meat weights vary seasonally, with the greatest meat weights during the 131 late spring and early summer. Meat weights drop considerably after the later 132 summer/early fall spawn and stay low until the spring. These patterns were 133 documented in NEFSC (2007) using observer data. Observers weigh scallop 134 meats in aggregate, so that it is not possible to distinguish which of the 135 shell height/meat weight parameters change seasonally. However, general 136 allometric principles suggest that most of the variation is in the intercept a137 rather than the slope (or power) parameter b. Haynes (1966) constructed a 138 number of monthly shell height/meat weight relationships, and did not find 139 any significant trend in the slopes. Thus, it was assumed in NEFSC (2007) 140 that all the seasonal variation in meat weights was due to variability in a. If 141 this is the case, seasonality would not affect the F_{MAX} reference point. For 142 this reason, seasonal variability was not considered a source of uncertainty 143 for this analysis. 144

Natural mortality M. Like most stocks, natural mortality is one of the 145 most uncertain parameters. However, dead "clapper" scallops (dead scallop 146 shells still attached at the hinge) are an indicator of recent natural mortality, 14 due to such causes as disease, high temperatures and sea star predation. The 148 clappers separate some time after death because of hinge degeneration. At 149 equilibrium, the rate of clappers being produced, ML, where L is the number 150 of live scallops, must equal the rate of loss of clappers C/S, where S is the 151 mean clapper separation time and C is the number of clappers. Solving this 152 for M gives: 153

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$$M = \frac{1}{S} \frac{C}{L} \tag{2}$$

so that natural mortality is proportional to the ratio of clappers to live scal-lops.

¹⁵⁷ Merrill and Posgay (1964) used this idea to estimate natural mortality. ¹⁵⁸ They estimated the clapper ratio C/L = 0.0662, and the mean separation

time S = 33 weeks = 33/52 years, to estimate an annual natural mortality 159 rate of $(52/33) * 0.0662 = 0.104 \approx 0.1$. Probably the greatest uncertainty in 160 this calculation is the mean separation time S. For example, Dickie (1955)161 estimated S to be 100 days (14.3 weeks). I assumed S was distributed as 162 a gamma random variable, with mean 33 weeks and standard deviation 15 163 weeks. The resulting distribution of M has the desirable characteristic of 164 being skewed to the right. This makes sense since, for example, a natural 165 mortality of M = 0.2 is possible, but an M = 0, or even close to zero, is not. 166 Note that because S appears in the denominator of (2), the mean value of 16 M is not equal to applying equation (2) with the mean value of S, so that 168 the original calculation of Merrill and Posgay (1964) was biased. 169

Fishery selectivity. Fishery selectivity s was estimated using an ascending
logistic curve of the form:

s

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$$=\frac{1}{1+\exp(\alpha-\beta H)}\tag{3}$$

where H is shell height. The mean, variances, and correlation of the α and β parameters were estimated based on CASA model runs from the last sea scallop assessment during the most recent time period. Note that fishery selectivity reflects targeting as well as gear selectivity.

Discard mortality. Sea scallops likely tolerate discarding fairly well, pro-177 vided they are returned to the water relatively promptly and they are not 178 damaged by the capture process or their time on deck. Further uncertainty 179 occurs in the summertime in the Mid-Atlantic, where summer SST exceeds 180 the thermal tolerance of sea scallops. Discard mortality was estimated at 181 20% in the last assessment, but there is little confidence in this number. 182 Here, discard mortality was simulated as a gamma distribution, with a mean 183 of 0.2 and a standard deviation of 0.15. 184

Incidental fishing mortality. Incidental fishing mortality occurs when scallops are killed but not captured by the gear. Let F_L be the landed fishing ¹⁸⁷ mortality rate and F_I be the rate of incidental fishing mortality. F_I should ¹⁸⁸ be proportional to F_L , say $F_I = iF_L$. Based on the studies of Caddy (1973) ¹⁸⁹ and Serchuk and Murawski (1989), *i* was estimated as 0.15 on Georges Bank ¹⁹⁰ and 0.04 in the Mid-Atlantic by NEFSC(2007). Because of the considerable ¹⁹¹ uncertainty in these numbers, *i* was simulated here with a gamma distribution ¹⁹² with these means and coefficients of variation of 0.75.

¹⁹³ Incorporating management uncertainty

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The actual fishing mortality realized may be different than the target fishing 194 mortality set by managers. Thus, for a fixed target fishing mortality F_{TARGET} , 195 the actual fishing mortality F_a is a random variable with density function 196 p(F). Denote by Y(F) the expected yield per recruit obtained by fishing at 197 F, and $Y_t(F)$ the expected yield per recruit obtained by setting the target 198 fishing mortality at F. Note that these will be different, even if the process of 199 setting the management targets is unbiased because yield per recruit curves 200 are non-linear. The expected yield per recruit obtained from setting the 201 target at F_{TARGET} is: 202

$$Y_t(F_{\text{TARGET}}) = \int_0^\infty p(F)Y(F) \, dF. \tag{4}$$

For these analyses, I assumed that the density function p(F) is normal (in principle, this needs to be truncated at 0, but in practice there is negligible probability that F < 0) with mean F_{TARGET} and standard deviation σ . The integral was estimated by discretization with a step size of 0.01.

It remains to estimate the standard deviation σ . The CASA stock assessment model generally estimates fishing mortalities with errors of between 0.01 to 0.02. However, these are estimates of past fishing mortalities, obtained when all the information is available. Managers set effort and/or quota levels based on forecasts that must contain more uncertainty than stock assessment estimates of prior years. The SAMS projection model used for forecasts in

the scallop fishery typically gives uncertainty in fishing mortalities of about 214 $\sigma = 0.04$ for short-term forecasts, based on bootstraps of initial conditions 215 and stochastic recruitment variability. This estimate does not include "model 216 error" such as uncertainties in model parameters or changes in fishing prac-21 tices. If this type of error is of similar magnitude and independent from the 218 stochastic error already quantified by the SAMS model, the total implemen-219 tation error is about $0.04\sqrt{2} \approx 0.06$. The analysis was conducted both with 220 $\sigma = 0.04$ as a lower bound and $\sigma = 0.06$. 22

222 Results and Discussion

The tradeoffs between probability of overfishing and losses in expected yield 223 are shown in Table 2 and Figure 1. Maximal expected yield per recruit 224 are obtained at somewhat higher (by about 0.07) fishing mortalities than 225 calculated in the last sea scallop assessment (NEFSC 2007). There are two 220 reasons for this. First, even though the Merrill and Posgay (1964) estimates 227 were used as the expected value of the clapper ratio and separation time 228 for the clappers, the mean natural mortality was about 0.13, rather than 229 the 0.1 estimated by Merrill and Posgay (1964), due to the uncertainty in 230 the denominator of equation (2). Secondly, the yield per recruit curve is 231 asymmetric, with a greater slope (in absolute magnitude) to the left of F_{MAX} 232 than to the right. As a result, expected yield per recruit will be optimized 233 by fishing at a level slightly greater than the point estimate of F_{MAX} . 234

Reducing fishing mortality near F_{MAX} produces considerable benefits (in terms of reduced risk of overfishing) at only a small cost (reduced expected yield per recruit). However, as fishing mortality is further reduced, benefits are reduced and costs increase. Basic cost/benefit theory states that the point of optimal cost/benefit will occur where the marginal benefit equals the marginal cost. The difficulty in applying this theory is that costs and

benefits are in incommensurate quantities, so that the value of a decreased 241 risk of overfishing compared to a loss in expected yield is subjective. Thus, 242 some judgement is required to decide the appropriate balance. The scallop 243 PDT suggested that the ABC fishing mortality should be set where the risk 244 of overfishing is 0.25, or where the loss of yield per recruit is 1% less than 245 optimal, whichever is less. According to Table 2, this would result in an ABC 246 fishing mortality target of 0.28. While this value is reasonable, arguments 24 can be made for just about any target between 0.2 and 0.3. 248

Performing similar analyses, but using target fishing mortality instead of actual fishing mortality, indicates that at lower fishing mortalities, implementation error increases both the risk of overfishing and the loss of yield per recruit due to precaution (Tables 3 and 4; Figures 2 and 3).

It is also of interest when setting the target is to calculate the probability of exceeding the ABC fishing mortality, since this triggers "accountability measures." Because implementation error is assumed to be normally distributed, this can be calculated simply from a table of (inverse) normal probabilities (Table 5).

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Table 1. Mean, standard error, and distributions of parameters used in the yield per recruit analysis.

Parameter	Purpose	Mean	S.E.	Corr.	Distribution
K	Growth	0.375	0.002	-0.6	Corr. Normal
L_{∞}	Growth	146.5	0.3	-0.6	Corr. Normal
a	$\mathrm{SH/MW}$	-10.70	0.27	-0.998	Corr. Normal
b	$\mathrm{SH/MW}$	2.942	0.055	-0.998	Corr. Normal
S	Nat. mort.	33/52 y	$15/52 { m y}$		Gamma
α	Selectivity	25.24	8.69	0.998	Corr. Normal
eta	Selectivity	0.23	0.08	0.998	Corr. Normal
F_D	Disc. mort.	0.2	0.15		Gamma
i	Incid. mort.	0.15	0.11		Gamma

(a) Georges Bank

(b) Mid-Atlantic

Parameter	Purpose	Mean	S.E.	Corr.	Distribution
K	Growth	0.495	0.004	-0.6	Corr. Normal
L_{∞}	Growth	131.6	0.4	-0.6	Corr. Normal
a	$\mathrm{SH/MW}$	-12.01	0.15	-0.997	Corr. Normal
b	$\mathrm{SH/MW}$	3.22	0.05	-0.997	Corr. Normal
S	Nat. mort.	33/52 y	$15/52 { m y}$		Gamma
α	Selectivity	21.67	2.77	0.998	Corr. Normal
β	Selectivity	0.214	0.03	0.998	Corr. Normal
F_D	Disc. mort.	0.2	0.15		Gamma
i	Incid. mort.	0.04	0.03		Gamma

Geo	rges Bank		Mid-Atlantic		;	Overall		
F	POF	%Loss	F	POF	%Loss	F	POF	% Loss
0.10	0	23	0.20	0.003	7.7	0.15	0	12.33
0.11	0	19.7	0.21	0.007	6.8	0.16	0	10.62
0.12	0	16.7	0.22	0.012	5.9	0.17	0.003	9.13
0.13	0	14.2	0.23	0.021	5.1	0.18	0.005	7.81
0.14	0	12.1	0.24	0.033	4.4	0.19	0.01	6.66
0.15	0.001	10.2	0.25	0.050	3.8	0.20	0.02	5.65
0.16	0.004	8.6	0.26	0.066	3.2	0.21	0.038	4.77
0.17	0.011	7.2	0.27	0.084	2.7	0.22	0.058	4
0.18	0.022	5.9	0.28	0.108	2.3	0.23	0.083	3.32
0.19	0.04	4.9	0.29	0.132	1.9	0.24	0.108	2.74
0.20	0.06	4	0.30	0.159	1.6	0.25	0.13	2.23
0.21	0.087	3.2	0.31	0.186	1.3	0.26	0.158	1.79
0.22	0.119	2.6	0.32	0.215	1.0	0.27	0.189	1.41
0.23	0.154	2	0.33	0.244	0.8	0.28	0.225	1.09
0.24	0.191	1.5	0.34	0.277	0.6	0.29	0.254	0.82
0.25	0.226	1.2	0.35	0.304	0.5	0.30	0.29	0.6
0.26	0.263	0.8	0.36	0.333	0.3	0.31	0.333	0.41
0.27	0.303	0.6	0.37	0.363	0.2	0.32	0.355	0.27
0.28	0.341	0.4	0.38	0.388	0.1	0.33	0.385	0.16
0.29	0.381	0.2	0.39	0.416	0.1	0.34	0.418	0.08
0.30	0.418	0.1	0.40	0.443	0.0	0.35	0.448	0.03
0.31	0.449	0	0.41	0.467	0.0	0.36	0.483	0
0.32	0.484	0	0.42	0.490	0.0	0.37	0.51	0
0.33	0.515	0	0.43	0.512	0.0	0.38	0.535	0.02
0.34	0.54	0	0.44	0.535	0.0	0.39	0.555	0.06
0.35	0.568	0.1	0.45	0.557	0.0	0.40	0.578	0.11

Table 2. Probability of overfishing (POF) and loss of yield per recruit (percetage loss compared to maximal) for sea scallops in Georges Bank, the Mid-Atlantic, and overall.

Table 3. Probability of overfishing (POF) and loss of yield per recruit (percetage loss compared to maximal) for sea scallops in Georges Bank, the Mid-Atlantic, and overall, with respect to target fishing mortality rates, assuming $\sigma = 0.04$ implementation uncertainty.

Georges Bank		Mid-Atlantic			Overall			
F_{target}	POF	%Loss	F_{target}	POF	% Loss	F_{target}	POF	% Loss
0.10	0.001	27.7	0.20	0.015	8.7	0.15	0.016	14.06
0.11	0.002	23.6	0.21	0.022	7.6	0.16	0.022	12.12
0.12	0.004	20.0	0.22	0.030	6.6	0.17	0.029	10.43
0.13	0.007	17.0	0.23	0.040	5.8	0.18	0.038	8.96
0.14	0.012	14.4	0.24	0.052	5.0	0.19	0.049	7.66
0.15	0.018	12.2	0.25	0.067	4.3	0.20	0.062	6.51
0.16	0.027	10.3	0.26	0.083	3.7	0.21	0.076	5.50
0.17	0.038	8.7	0.27	0.102	3.2	0.22	0.093	4.63
0.18	0.053	7.3	0.28	0.122	2.7	0.23	0.111	3.86
0.19	0.070	6.1	0.29	0.145	2.3	0.24	0.131	3.20
0.20	0.091	5.0	0.30	0.169	1.9	0.25	0.153	2.62
0.21	0.114	4.1	0.31	0.194	1.6	0.26	0.177	2.12
0.22	0.141	3.4	0.32	0.220	1.3	0.27	0.201	1.69
0.23	0.170	2.7	0.33	0.247	1.1	0.28	0.227	1.33
0.24	0.201	2.2	0.34	0.275	0.9	0.29	0.254	1.02
0.25	0.234	1.7	0.35	0.302	0.7	0.30	0.281	0.76
0.26	0.268	1.3	0.36	0.330	0.5	0.31	0.309	0.55
0.27	0.303	1.0	0.37	0.357	0.4	0.32	0.337	0.37
0.28	0.337	0.8	0.38	0.384	0.3	0.33	0.364	0.24
0.29	0.372	0.6	0.39	0.410	0.2	0.34	0.392	0.14
0.30	0.406	0.4	0.40	0.435	0.2	0.35	0.419	0.06
0.31	0.439	0.3	0.41	0.460	0.1	0.36	0.445	0.02
0.32	0.471	0.3	0.42	0.484	0.1	0.37	0.470	0.00
0.33	0.501	0.2	0.43	0.507	0.1	0.38	0.495	0.00
0.34	0.530	0.2	0.44	0.529	0.0	0.39	0.518	0.03
0.35	0.558	0.3	0.45	0.551	0.0	0.40	0.541	0.07

Table 4. Probability of overfishing (POF) and loss of yield per recruit (percetage loss compared to maximal) for sea scallops in Georges Bank, the Mid-Atlantic, and overall, with respect to target fishing mortality rates, assuming $\sigma = 0.06$ implementation uncertainty.

Georges Bank		Mid-Atlantic			Overall			
F_{target}	POF	%Loss	F_{target}	POF	% Loss	F_{target}	POF	%Loss
0.1	0.006	27.5	0.2	0.033	9.4	0.15	0.016	16.22
0.11	0.009	25.2	0.21	0.042	8.2	0.16	0.022	13.71
0.12	0.014	22.9	0.22	0.053	7.2	0.17	0.029	11.77
0.13	0.019	20.5	0.23	0.064	6.3	0.18	0.038	10.09
0.14	0.026	19.4	0.24	0.078	5.4	0.19	0.049	8.63
0.15	0.034	16.7	0.25	0.093	4.7	0.2	0.062	7.36
0.16	0.044	14.1	0.26	0.110	4.1	0.21	0.076	6.25
0.17	0.057	11.9	0.27	0.129	3.5	0.22	0.093	5.28
0.18	0.071	10.1	0.28	0.149	3.0	0.23	0.111	4.42
0.19	0.088	8.5	0.29	0.170	2.5	0.24	0.131	3.67
0.2	0.107	7.1	0.3	0.192	2.2	0.25	0.153	3.03
0.21	0.128	5.9	0.31	0.216	1.8	0.26	0.177	2.47
0.22	0.151	4.9	0.32	0.240	1.5	0.27	0.201	1.99
0.23	0.176	4.1	0.33	0.265	1.3	0.28	0.227	1.58
0.24	0.202	3.3	0.34	0.290	1.0	0.29	0.254	1.23
0.25	0.231	2.7	0.35	0.316	0.9	0.3	0.281	0.93
0.26	0.260	2.2	0.36	0.341	0.7	0.31	0.309	0.68
0.27	0.291	1.8	0.37	0.367	0.6	0.32	0.337	0.48
0.28	0.321	1.4	0.38	0.392	0.4	0.33	0.365	0.32
0.29	0.353	1.2	0.39	0.417	0.3	0.34	0.392	0.20
0.3	0.384	0.9	0.4	0.441	0.3	0.35	0.419	0.11
0.31	0.415	0.8	0.41	0.465	0.2	0.36	0.445	0.05
0.32	0.445	0.6	0.42	0.489	0.2	0.37	0.470	0.01
0.33	0.475	0.5	0.43	0.511	0.1	0.38	0.495	0.00
0.34	0.504	0.4	0.44	0.533	0.1	0.39	0.519	0.01
0.35	0.532	0.4	0.45	0.554	0.1	0.4	0.541	0.04

Reduction	$P(F > F_{\rm ABC})$	$P(F > F_{\rm ABC})$
in F	$\sigma = 0.04$	$\sigma = 0.06$
0.01	0.401	0.434
0.02	0.309	0.369
0.03	0.227	0.309
0.04	0.159	0.252
0.05	0.106	0.202
0.06	0.067	0.159
0.07	0.040	0.122
0.08	0.023	0.091
0.09	0.012	0.067
0.10	0.006	0.048
0.11	0.003	0.033
0.12	0.001	0.023

Table 5. Risk of exceeding the ABC and hence encountering accountability measures at various reductions in target fishing mortalities below the ABC fishing mortality.

Figure legends

Figure 1. The probability of overfishing (solid) and loss in yield per recruit (dashed) for (a) Georges Bank, (b) Mid-Atlantic and (c) overall, as a function of true fishing mortality.

Figure 2. Figure 1. The probability of overfishing (solid) and loss in yield per recruit (dashed) for (a) Georges Bank, (b) Mid-Atlantic and (c) overall, as a function of target fishing mortality with implementation error $\sigma = 0.04$.

Figure 3. The probability of overfishing (solid) and loss in yield per recruit (dashed) for (a) Georges Bank, (b) Mid-Atlantic and (c) overall, as a function of target fishing mortality with implementation error $\sigma = 0.06$.

Figure 1 (a) (b) (c) 0.6 0.6 0.6 Prob OF Frac loss YPR 0.5 Prob OF Frac loss YPR 0.5 Prob OF Frac loss YPR 0.5 -0.4 -0.4 0.4 -0.3 0.3 -0.3 -0.2 -0.2 0.2 -0.1 0.1 0.1 0.0 0.0 0.0 0.15 0.20 0.25 0.30 0.35 0.10 0.15 0.20 0.25 0.30 0.35 0.40 0.45 0.10 0.15 0.20 0.25 0.30 0.35 0.40 Fishing mortality Fishing mortality Fishing mortality Figure 2 (a) (b) (c) 0.6 0.6 0.6 Prob OF Frac loss YPR 0.5 Prob OF Frac loss YPR Prob OF Frac loss YPR 0.5 0.5 0.4 -0.4 · 0.4 -

0.3

0.2 -

0.1 -

0.0

0.10 0.15 0.20

0.25

Target fishing mortality

0.30

0.35 0.40

0.45

0.30

0.35

0.3 -

0.2

0.1 -

0.0 -

0.10

0.15

0.25

Target fishing mortality

0.20

0.30

0.40

0.35

0.3 -

0.2

0.1 -

0.0

0.15

0.20

Target fishing mortality

0.25

